

Summary of Lecture 6

In which we study the spectrum of random graphs.

In past lectures we used the fact that, for every $p(n) > \frac{\log n}{n}$, there is a high probability that if we sample $G \sim G_{n,p(n)}$ and let A be its adjacency matrix, we have $\|A - pJ\| \leq O(\sqrt{p(n) \cdot n})$. In particular, there is a high probability that $\|A - J/2\| \leq O(\sqrt{n})$ when A is the adjacency matrix of a graph sampled from $G_{n, \frac{1}{2}}$.

Today we will prove a slightly weaker result, and give an overview of the proofs of the stronger results.

We will use the following basic facts from linear algebra. Let M be a real symmetric matrix: then M has real eigenvalues, and let us call them $\lambda_1, \dots, \lambda_n$ so that $|\lambda_1| \geq |\lambda_2| \geq \dots \geq |\lambda_n|$. By definition, $\|M\| = |\lambda_1|$. We have

1. $\sum_i M_{ii} = \sum_i \lambda_i$, and this quantity is called the *trace* of the matrix.
2. The eigenvalues of M^k are $\lambda_1^k, \dots, \lambda_n^k$

From the above two facts we have that, if k is even,

$$\sum_i (M^k)_{ii} = \sum_i \lambda_i^k = \sum_i |\lambda_i|^k$$

and so

$$\|M\|^k \leq \sum_i (M^k)_{ii} \leq n \cdot \|M\|^k$$

If we choose $k = \ln n$, we have

$$\|M\| \leq \left(\sum_i (M^k)_{ii} \right)^{\frac{1}{k}} \leq e \cdot \|M\|$$

Thus we can approximate the spectral norm of a matrix by a computation of the trace of a power of the matrix. If M is a random matrix, then computing $\mathbb{E} \sum_i (M^k)_{ii}$ will give us high-probability upper bounds to the spectral norm.

Let A be the adjacency matrix of a $G_{n, 1/2}$ random graph; we let M be the random matrix $M := 2 \cdot (A - \mathbb{E} A)$. We see that M is such that the diagonal entries are zero, all other entries

are equally likely to be $+1$ or -1 , and, other than satisfying the constraints $M_{ij} = M_{ji}$, all entries are mutually independent.

We first notice that, by symmetry

$$\mathbb{E} \sum_i (M^k)_{ii} = n \cdot \mathbb{E}(M^k)_{1,1}$$

and by definition of matrix product and by linearity of expectation we have

$$\mathbb{E}(M^k)_{1,1} = \sum_{a_1, \dots, a_{k-1} \in V} \mathbb{E} M_{1,a_1} \cdot M_{a_1,a_2} \cdots M_{a_{k-2},a_{k-1}} \cdot M_{a_{k-1},1}$$

For each entry (a, b) , we have that $\mathbb{E} M_{a,b}^t = 0$ if t is odd and $\mathbb{E} M_{a,b}^t = 1$ if t is even, so the expressions whose average we want to compute above have average zero if there is a $\{a, b\}$ that occurs an odd number of times in the sequence $\{1, a_1\}, \dots, \{a_{k-1}, 1\}$, and the average is one if all pairs occur an even number of times in the sequence.

Thus, $\mathbb{E}(M^k)_{1,1}$ is equal to the number of sequences $1, a_1, \dots, a_{k-1}, 1$ such that, in the sequence $\{1, a_1\}, \dots, \{a_{k-1}, 1\}$, each pair appears an even number of times.

We prove that such number is at most $2^k \cdot k^{k/2} \cdot n^{1+k/2}$ by showing how to represent each sequence using at most $k + \frac{k}{2} \log k + \frac{k}{2} \log n$ in the following way: each element a_i is represented either as $(0, a_i)$, which takes $1 + \log n$ bits, if a_i appears for the first time in the sequence at location i , and as $(0, j)$ otherwise, where $j < i$ is such that $a_i = a_j$, which requires $1 + \log k$ bits. Using the fact that the sequence can contain at most $k/2$ distinct pairs and hence at most $k/2$ distinct vertices (not counting the start) we get the bound.

This means that

$$\mathbb{E} \sum_i (M^k)_{ii} \leq 2^k \cdot k^{k/2} \cdot n^{1+k/2}$$

If we take $k = \lg n$ we have

$$\mathbb{P}[\|M\| > t] = \mathbb{P}[\|M\|^k > t^k] \leq \frac{\mathbb{E} \sum_i (M^k)_{ii}}{t^k} \leq \left(\frac{2e\sqrt{nk}}{t} \right)^k$$

And so with high probability $\|M\| \leq 4e\sqrt{n \ln n}$.

A more careful counting shows that the number of sequences is actually at most $2^{O(k)} n^{k/2}$, leading to the conclusion that $\|M\| \leq O(\sqrt{n})$ with high probability. We can understand the set of sequences that we want to count in the following way: a sequence $1, a_1, \dots, a_{k-1}, 1$ defines a graph over the vertices $\{1, a_1, \dots, a_{k-1}, 1\}$ with edges $\{\{1, a_1\}, \dots, \{a_{k-1}, 1\}\}$. Because of repetitions, we have at most $k/2$ edges, and hence at most $1 + k/2$ vertices. In our previous counting argument, we upper bounded the number of sequences with ℓ distinct vertices (not counting the start vertex 1) as $\binom{k}{\ell} \cdot k^{k-\ell} \cdot n^\ell$, which increases with ℓ and maxes out at $\ell = k/2$, thus it is interesting to see, as a first step toward a better proof, if the

count can be improved in the $\ell = k/2$. In this case, the graph is a tree, which is visited in a depth-first search order. This means that, at every step in the sequence, we either see a new vertex or we *backtrack* in our depth-first search, and we only need to spend one bit to represent the latter even, leading to a much better $2^k \cdot n^{k/2}$ bound. Obtaining a $2^{O(k)} n^{k/2}$ bound for all sequences is much harder and we will not attempt to do so.

Let us briefly review how to generalize the analysis to $G_{n,p}$ for $p < 1/2$. In this case we will look at

$$M := A - \mathbb{E}A = A - pJ + pI$$

and we see that M is symmetric, it has zeroes on the diagonals, and the off-diagonal entries satisfy

$$\begin{aligned} \mathbb{E} M_{i,j} &= 0 \\ \mathbb{E} M_{i,j}^2 &= p - p^2 \leq p \\ \mathbb{E} M_{i,j}^k &\leq p \quad \forall k > 2 \end{aligned}$$

As before

$$\mathbb{E} \text{trace}(M^k) = n \mathbb{E}(M^k)_{1,1}$$

and

$$\mathbb{E}(M^k)_{1,1} = \sum_{a_1, \dots, a_{k-1} \in V} \mathbb{E} M_{1,a_1} \cdot M_{a_1,a_2} \cdots M_{a_{k-2},a_{k-1}} \cdot M_{a_{k-1},1}$$

Now the expectation on the right-hand side is zero if there is a pair that occurs only once in the sequence. Otherwise, it is p^ℓ where ℓ is the number of distinct pairs in the sequence. The contribution to the weighted sum of sequences with $k/2$ distinct “edges” is thus $2^{O(k)} p^{k/2} n^{k/2}$. When $p > \frac{\log n}{n}$, this is the dominant term, and the whole weighted sum is at most $2^{O(k)} p^{k/2} n^{k/2}$, leading to a high probability bound $\|M\| \leq O(\sqrt{pn})$.

Unfortunately, for $p = o\left(\frac{\log n}{n}\right)$ we do not have the high probability bound $\|M\| \leq O(\sqrt{pn})$. To see why not, first observe that, for every $p > 1/n$ there is a high probability that the *largest* degree in the graph is $\Omega\left(\frac{\log n}{\log \log n}\right)$. Furthermore, if a graph has a vertex of large degree, then $\|A - pJ\|$ will have large spectral norm.

Lemma 1 *Let G be a graph, A its adjacency matrix, and suppose that G has a vertex v of degree $D \geq 1$. Then, for every p , we have*

$$\|A - pJ\| \geq \sqrt{D} - 2pD$$

PROOF: Let u_1, \dots, u_D be the D neighbors of v . Consider the vector \mathbf{x} defined as $\mathbf{x}_v = 1$, $\mathbf{x}_{u_i} = 1/\sqrt{D}$ for $i = 1, \dots, D$, and $\mathbf{x}_z = 0$ for all other vertices z .

Then

$$\begin{aligned}\|\mathbf{x}\|^2 &= 2 \\ \mathbf{x}^T A \mathbf{x} &\geq 2\sqrt{D} \\ \mathbf{x}^T pJ \mathbf{x} &= p \cdot \left(\sum_z \mathbf{x}_z \right)^2 \leq p \cdot (1 + \sqrt{D})^2 \leq 4pD\end{aligned}$$

and

$$\|A - pJ\| \geq \frac{\mathbf{x}^T (A - pJ) \mathbf{x}}{\|\mathbf{x}\|^2} \geq \sqrt{D} - 2pD$$

□

Thus, for $p < \frac{\log n}{n}$, we have with high probability $\|A - pJ\| \geq \Omega\left(\sqrt{\frac{\log n}{\log \log n}}\right)$